STA 331 2.0 Stochastic Processes

8. Birth and Death Processes

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Birth and Death Processes

- The birth-and-death process is a subclass of continuous-time Markov chains.
- The birth-and-death processes are characterized by the property that whenever a transition occurs from one state to another, then this transition can be to a neighbouring state only.

Transition types

- a transition occurs from one state to another and this transition can be to a neighbouring state only.
 - Eg: State space $S = \{0, 1, 2, ..., i, ...\}$
 - transition that occurs from state i, can be only to a neighboring state (i-1) or (i+1).

Birth rate and Death rate

Birth rate

 λ_i - birth rate from state i, $i \ge 0$

Death rate

 μ_i - death rate from state i, $i \geq 0$

Queueing systems

- 1. Birth equivalent to the arrival of a customer.
- 2. Death equivalent to the departure of a served customer.

Notations

A continuous-time Markov chain $[X(t), t \in T]$ with state space $S = \{0, 1, 2, ...\}$ with rates

$$q_{i,i+1} = \lambda_i, i = 0, 1, ...,$$

$$q_{i,i-1} = \mu_i, i = 1, 2, ...,$$

$$q_{i,j} = 0, \ j \neq i \pm 1, \ j \neq i, \ i = 0, 1, ..., \ \text{and}$$

$$q_i = (\lambda_i + \mu_i), i = 0, 1, ..., \text{ and } \mu_0 = 0.$$

Pure birth process, pure death process, birth-and-death process

- i) a pure birth process if $\mu_i = 0$ for i = 1, 2, ...
 - No decrements, only increments.
- ii) a pure death process if $\lambda_i = 0$ for i = 1, 2, ...
 - No increments, only decrements.
- iii) a birth-and-death process if some of the λ_i 's and some of the μ_i 's are positive.

Examples of random phenomena modelled through birth and death processes

- Spread of epidemic disease
- Mutant gene dynamics
- Cell kinetics (proliferation of cancer cells)

Special cases

- 1. Linear birth process: Yule-Furry process
- 2. Linear death process
- 3. Linear birth and death process
- 4. M/M/I queue

Pure Birth Process

- Special case of a continuous-time Markov process and a generalisation of a Poisson process.
- Consider a population of individuals where only the appearances of new individuals, which are called "birth" occur.

General birth processes

Let us consider a birth process whose total number of individuals at time t is denoted by a discrete random variable N(t). As parameter t varies $\{N(t): t \geq 0\}$ represents a stochastic process with a continuous parameter (time) space and a discrete state space.

Let us assume that the birth rate depends on the present size of the population. Further we assume that the births occur according to the following postulates:

$$P[N(t+h) = n+k|N(t) = n] = \begin{cases} \lambda_n h + o(h), & k=1\\ o(h), & k \geq 2\\ 1 - \lambda_n h + o(h), & k = 0 \end{cases}$$

General birth processes (cont)

Condition 1

$$P[N(t+h) = n+k|N(t) = n] = \begin{cases} \lambda_n h + o(h), & k=1\\ o(h), & k \geq 2\\ 1 - \lambda_n h + o(h), & k = 0 \end{cases}$$

where λ_n is the rate at which the births occur at time t and n being the size of the population at time t.

Condition 2

Your turn

Compare the differences in conditions between Poisson process, Non-homogeneous Poisson Process and Birth Process

Goal: Probability Mass Function of N(t)

What is the probability that the population size at a given time, t, equals N(t) = n?

$$P_n(t) = P[N(t) = n] = ?$$

For example,

$$P_0(t) = P[N(t) = 0] = ?$$

$$P_1(t) = P[N(t) = 1] = ?$$

$$P_2(t) = P[N(t) = 2] = ?$$

Linear Birth Process (Yule-Furry Process)

When, $\lambda_n = n\lambda$, i.e. when the birth rate is linear in the present size of the population.

Then the pure birth process is said to a **Linear Birth Process** or **Yule-Furry Process**.

Let is assume that there is only one individual in the population initially, N(0) = 1. It can be shown that for any t > 0.

$$P(N(t)=0)=0$$

$$P(N(t) = n) = e^{-\lambda t} (1 - e^{-\lambda t})^{n-1}, n \ge 1.$$

Proof (general situation):

For
$$n = 0$$

$$P_0(t+h) = P(N(t) = 0)P(N(t+h) = 0|N(t) = 0)$$

$$P_0(t+h) = P_0(t)(1 - \lambda_0 h + o(h))$$

i.e.

$$P_0(t+h) = P_0(t) - \lambda_0 h P_0(t) + o(h) P_0(t)$$

$$\lim_{h \to 0} \frac{P_0(t+h) - P_0(t)}{h} = -\lim_{h \to 0} \lambda_0 P_0(t) + \lim_{h \to 0} \frac{o(h)}{h} P_0(t)$$

i.e.

$$P_0'(t) = -\lambda_0 P_0(t).$$

We assume that there is only one individual in the population initially, N(0) = 1. Hence, P[N(t) = 0] = 0.

That is $P_0(t) = 0$.

Proof: (cont)

For n > 1

$$P_n(t+h) = P(N(t) = n)P(N(t+h) = n|N(t) = n) + P(N(t) = n-1)P(N(t+h) = n|N(t) = n-1) + \sum_{r=2}^{n-1} P(N(t) = n-r)P(N(t+h) = n|N(t) = n-r)$$

i.e

$$P_n(t+h) = P_n(t)(1 - \lambda_n h + o(h)) + P_{n-1}(t)(\lambda_{n-1} h + o(h)) + o(h)$$

Proof: (cont)

$$P_{n}(t+h) = P_{n}(t) - \lambda_{n}hP_{n}(t) + \lambda_{n-1}hP_{n-1}(t) + o(h) \text{ for } n \ge 1$$

$$\lim_{h \to 0} \frac{P_{n}(t+h) - P_{n}(t)}{h} = -\lambda_{n}P_{n}(t) + \lambda_{n-1}P_{n-1}(t) + \lim_{h \to 0} \frac{o(h)}{h}$$
i.e.

$$P_n'(t) = -\lambda_n P_n(t) + \lambda_{n-1} P_{n-1}(t)$$
 for $n \ge 1$.

Therefore the partial differential-difference equations is

For
$$n \ge 1$$
, $P'_n(t) = -\lambda_n P_n(t) + \lambda_{n-1} P_{n-1}(t)$.

When n = 1

$$P_1'(t) = -\lambda_1 P_1(t),$$

$$\int \frac{P_1'(t)}{P_1(t)} dt = -\lambda_1 \int dt,$$

$$InP_1(t) = -\lambda_1 t + c$$
 $P_1(t) = c_1 e^{-\lambda_1 t}$

When t = 0, $c_1 = 1$

$$P_1(t) = e^{-\lambda_1 t}$$

When n = 2

$$P_2'(t) = -\lambda_2 P_2(t) + \lambda_1 P_1(t),$$

$$P_2'(t) + \lambda_2 P_2(t) = \lambda_1 e^{-\lambda_1 t},$$

Multiply by $e^{\lambda_2 t}$

$$egin{align} P_2'(t)e^{\lambda_2 t} + \lambda_2 P_2(t)e^{\lambda_2 t} &= \lambda_1 e^{-\lambda_1 t}e^{\lambda_2 t}, \ &\int rac{d}{dt}[e^{\lambda_2 t}P_2(t)]dt = \int \lambda_1 e^{(\lambda_2 - \lambda_1)t}dt, \ &e^{\lambda_2 t}P_2(t) &= rac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= rac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= rac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= rac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} + c \ &e^{\lambda_2 t}P_2(t) &= \frac{\lambda_1 e^{(\lambda_2 - \lambda_1)t}}{\lambda_2 - \lambda_1} +$$

When t = 0,

We know that $P_2(0) = 0$. hence,

$$c = -\frac{\lambda_1}{\lambda_2 - \lambda_1}$$
.

Hence,

$$P_2(t) = \frac{\lambda_1}{\lambda_2 - \lambda_1} e^{-\lambda_2 t} [e^{(\lambda_2 - \lambda_1)t} - 1]$$

When,

$$\lambda_n = n\lambda$$
.

That is the birth rate is linear in the present size of the population.

Let us assume that there is only one individual in the population initially. That is N(0) = 1.

Then the difference-differential equations of the linear birth process takes the form

$$P_n'(t)=-n\lambda P_n(t)+(n-1)\lambda P_{n-1}(t)$$
 for $n\geq 1$ with the initial conditions $P_1(0)=1$ and $P_n(0)=0$ for $n\geq 2$.

 $P_n'(t) = -n\lambda P_n(t) + (n-1)\lambda P_{n-1}(t)$ for $n \ge 1$ with the initial conditions $P_1(0) = 1$ and $P_n(0) = 0$ for $n \ge 2$.

Multiplying the equation for n by z^n and summing over all n we obtain

$$\frac{\partial}{\partial t} \sum_{n=1}^{\infty} P_n(t) z^n = -\lambda z \frac{\partial}{\partial z} \sum_{n=1}^{\infty} P_n(t) z^n + \lambda z^2 \frac{\partial}{\partial z} \sum_{n=1}^{\infty} P_{n-1}(t) z^{n-1}$$

Let $\prod(z,t) = \sum_{n=1}^{\infty} P_n(t)z^n$. Then the above equations becomes

$$\frac{\partial \prod(z,t)}{\partial t} = -\lambda z \frac{\partial \prod(z,t)}{\partial z} + \lambda z^2 \frac{\partial \prod(z,t)}{\partial z}$$

i.e.
$$\frac{\partial \prod(z,t)}{\partial t} = \lambda z (z-1) \frac{\partial \prod(z,t)}{\partial z}$$

 $\frac{\partial \prod(z,t)}{\partial t} - \lambda z (z-1) \frac{\partial \prod(z,t)}{\partial z} = 0$

Subsidiary equations take the form

$$\frac{dt}{1} = \frac{dz}{-\lambda z(z-1)} = \frac{d\prod}{0}$$

Two independent solutions can be obtained one from $d \prod = 0$ and the other from $-\lambda dt = \frac{dz}{z(z-1)}$.

$$d \prod = 0 \Rightarrow \prod (z, t) = constant.$$

$$-\lambda dt = rac{dz}{z(z-1)} \Rightarrow rac{z}{z-1}e^{-\lambda t} = constant.$$

The general solution can be written as

$$\prod(z,t)=f\Bigl(rac{z}{z-1}e^{-\lambda t}\Bigr)$$
 where f is an arbitrary function.

The initial conditions $P_1(0)=1$ and $P_n(0)=0$ for $n\geq 2$ imply that $\prod (z,0)=z$.

$$\therefore \prod (z,0) = f\left(\frac{z}{z-1}\right) = z.$$

Let $\omega=\frac{\mathbf{z}}{\mathbf{z}-1}\Rightarrow \mathbf{z}=\frac{\omega}{\omega-1}$ and hence we obtain $f(\omega)=\frac{\omega}{\omega-1}$.

$$\therefore \prod (z,t) = \frac{\frac{z}{z-1}e^{-\lambda t}}{\frac{z}{z-1}e^{-\lambda t} - 1} = \frac{ze^{-\lambda t}}{ze^{-\lambda t} - (z-1)} = \left(1 - \frac{z-1}{z}e^{-\lambda t}\right)^{-1}$$

Considering coefficients of z^n we have

$$P_n(t) = e^{-\lambda t} (1 - e^{-\lambda t})^{n-1} \text{ for } n \ge 1.$$

In proving the above results we assume that initially there is only one individual in the population. That is N(0)=1.

Now let's prove for the case $N(0) = a, a \ge 1$. For that we use moment generating functions.

Moment generating function of N(t)

Let

$$M_{N(t)}(\theta, t) = E[e^{N(t)\theta}],$$

be the moment generating function of N(t). Then, for t > 0,

$$egin{align} M_{N(t)}(heta,t) &= \sum_{n=0}^{\infty} e^{n heta} P(N(t)=n) \ &= \sum_{n=0}^{\infty} e^{n heta} P_n(t). \end{split}$$

We assume that N(0) = a > 0. Hence, $P_n(t) = 0$ for all n < a. Hence,

$$M_{N(t)}(\theta,t) = \sum_{n=a}^{\infty} e^{n\theta} P_n(t).$$
 (2)

Now we take derivative w.r.t θ . Then we get,

$$\frac{\partial}{\partial \theta} M_{N(t)}(\theta, t) = \sum_{n=a}^{\infty} n e^{n\theta} P_n(t).$$

The derivative w.r.t t is

$$\begin{split} \frac{\partial}{\partial t} M_{N(t)}(\theta, t) &= \sum_{n=a}^{\infty} e^{n\theta} P'_{n}(t) \\ &= \sum_{n=a}^{\infty} e^{n\theta} [-n\lambda P_{n}(t) + (n-1)\lambda P_{n-1}(t)] \\ &= -\sum_{n=a}^{\infty} n e^{n\theta} \lambda P_{n}(t) + \sum_{n=a}^{\infty} (n-1)e^{n\theta} \lambda P_{n-1}(t) \end{split}$$

3) ₂

Since $P_{a-1}(t) = 0$, the second summation starts at a + 1. Hence.

$$\begin{split} \frac{\partial}{\partial t} M_{N(t)}(\theta, t) &= -\sum_{n=a}^{\infty} n e^{n\theta} \lambda P_n(t) + \sum_{n=a+1}^{\infty} (n-1) e^{n\theta} \lambda P_{n-1}(t) \\ &= -\sum_{n=a}^{\infty} n e^{n\theta} \lambda P_n(t) + \sum_{m=a}^{\infty} m e^{(m+1)\theta} \lambda P_m(t) \\ &= -\lambda \sum_{n=a}^{\infty} n e^{n\theta} P_n(t) + \lambda e^{\theta} \sum_{m=a}^{\infty} m e^{m\theta} P_m(t) \\ &= -\lambda \frac{\partial}{\partial \theta} M_{N(t)}(\theta, t) + \lambda e^{\theta} \frac{\partial}{\partial \theta} M_{N(t)}(\theta, t) \\ &= \lambda (e^{\theta} - 1) \frac{\partial}{\partial \theta} M_{N(t)}(\theta, t) \end{split}$$

$$\frac{\partial}{\partial t} M_{N(t)}(\theta, t) - \lambda (e^{\theta} - 1) \frac{\partial}{\partial \theta} M_{N(t)}(\theta, t) = 0.$$
 (5)

Note:

A partial differential equation (PDE) for a function z(x, y) is Lagrange type if it takes the form (General form of first-order quasilinear PDE)

$$P(x, y, z) \frac{\partial z}{\partial x} + Q(x, y, z) \frac{\partial z}{\partial y} = R(x, y, z).$$
 (6)

The associated characteristic system of ordinary differential equations.

Note (cont)

$$\frac{dx}{P(x,y,z)} = \frac{dy}{Q(x,y,z)} = \frac{dz}{R(x,y,z)}.$$
 (7)

is known as the characteristic (auxiliary) system of equation (5). Suppose that two independent particular solutions of this system have been found in the form

 $u(x, y, z) = C_1$ and $v(x, y, z) = C_2$, where where C_1 and C_2 are arbitrary constants.

Then the general solution to equation (5) can be written as

$$\phi(u,v)=0 \tag{8}$$

where ϕ is an arbitrary function of two variables.

Note (cont.)

With equation (6) solved for v, one often specifies the general solution in the form $v=\psi(u)$, where $\psi(u)$ is an arbitrary function of one variable. The ψ can be determined using the boundary conditions.

Revisit equation 4,

$$\frac{\partial}{\partial \theta} M_{N(t)}(\theta, t) - \lambda (e^{\theta} - 1) \frac{\partial}{\partial \theta} M_{N(t)}(\theta, t) = 0.$$
 (9)

According to the auxiliary system of equation in (6),

$$\frac{dt}{1} = \frac{d\theta}{-\lambda(e^{\theta} - 1)} = \frac{M_{N(t)}}{0}$$
$$\frac{dt}{1} = \frac{dM_{N(t)}}{0}$$

$$rac{dM_{N(t)}}{dt}=0 \Rightarrow M_{N(t)}(heta,t)=constant.$$

Furthermore consider,

$$\frac{dt}{1} = \frac{d\theta}{-\lambda(e^{\theta} - 1)}$$

$$\lambda dt = -\frac{1}{(e^{\theta} - 1)} d\theta$$

$$= \frac{-e^{-\theta}}{1 - e^{-\theta}} d\theta$$
(10)

From equation (9) we can write

$$\lambda t = -\ln(1 - e^{-\theta}) + c$$

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Furthermore

$$ln(e^{\lambda t}) + ln(1 - e^{-\theta}) = c.$$

Hence,

$$e^{\lambda t}(1-e^{-\theta})=constant.$$

Hence, the general solution for eq(8) is

$$M_{N(t)}(\theta, t) = \Psi[e^{\lambda t}(1 - e^{-\theta})].$$

The boundary conditions $P_a(0)=1$, and $P_n(0)$ for $n\neq a$, imply that $M_{N(t)}(\theta,0)=\sum_{n=a}^{\infty}e^{n\theta}P_n(0)=e^{a\theta}$,

$$M_{N(t)}(\theta,0)=e^{a\theta}=\Psi(1-e^{-\theta}).$$

Let
$$\alpha = 1 - e^{-\theta}$$
. Then, $e^{\theta} = (1 - \alpha)^{-1}$. Hence,

$$e^{a\theta} = \Psi(\alpha) = (1 - \alpha)^{-a}$$
.

Therefore,

$$M_{N(t)}(\theta, t) = \Psi[e^{\lambda t}(1 - e^{-\theta})] = [1 - e^{\lambda t}(1 - e^{-\theta})]^{-a}.$$

Let $p = e^{-\lambda t}$ and p + q = 1. Then,

$$M_{N(t)}(\theta, t) = [1 - p^{-1}(1 - e^{-\theta})]^{-a} = \left[\frac{p - 1 + e^{-\theta}}{p}\right]^{-a} = \left(\frac{p}{e^{-\theta} - q}\right)^{a}$$

Now from this MGF, we can derive the moments of N(t).

It can be shown that

$$E(N(t))=a/p=ae^{\lambda t}$$
 and $V[N(t)]=a(1-p)/p^2=a(1-e^{-\lambda t})e^{2\lambda t}.$

Furthermore, we recognize the above MGF is in the form of the MGF of a negative binomial random variable Y, with probability mass function

$$P(Y = y) = {}^{y-1} C_{a-1}p^{a-1}q^{y-1-(a-1)}p = {}^{y-1} C_{a-1}p^aq^{y-a}$$
, for $y = a, a+1, ...$

Hence,

$$P(N(t) = n) = {}^{n-1} C_{a-1}p^aq^{n-a} = {}^{n-1} C_{a-1}e^{-\lambda ta}(1 - e^{-\lambda t})^{n-a}$$
 for $n = a, a + 1, ...$

Moment generating function of N(t) - Alternative approach

Using probability generating functions. Let $G_{N(t)}(s, t)$ is called the probability generating function,

$$G_{N(t)}(s,t) = E(s^{N(t)}) = \sum_{n=0}^{\infty} s^n P_n(t).$$

The coefficients of s^n of the expansion of $G_{N(t)}(s,t)$ will give $P_n(t)$.

Summary:

When,
$$N(0) = 1$$

$$P(N(t)=0)=0$$

$$P(N(t) = n) = e^{-\lambda t} (1 - e^{-\lambda t})^{n-1}, n \ge 1.$$

When,
$$N(0) = a$$

$$P(N(t) = n) = {}^{n-1} C_{a-1}p^aq^{n-a} = {}^{n-1} C_{a-1}e^{-\lambda ta}(1 - e^{-\lambda t})^{n-a}$$
 for $n = a, a + 1, ...$

Exercise

Consider a pure birth process on the states $\{0, 1, ..., N\}$ for which $\lambda_k = (N - k)\lambda$ for k = 0, 1, ..., N. Suppose N(0) = 0. Find Pn(t) = P(X(t) = n) for n = 0, 1 and 2.